> Robert C. Vaughan

Motivation

Zero Densit Estimates

Counting Zeros

Background to Guth-Maynard

Robert C. Vaughan

September 4, 2024

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Motivation

Zero Densit Estimates

Counting Zeros • The question is, how do we deal with $\pi(x+h) - \pi(x),$ or equivalently

$$\psi(x+h)-\psi(x)$$

when

$$h = x^{\theta}$$
 with $\theta < 1$.

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$$\pi(x+h)-\pi(x),$$

or equivalently

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 with $\theta < 1$.

• The classical prime number theorem is not good enough, so try the explicit formula

$$\psi_0(x) = x - \sum_{\rho} \frac{x^{
ho}}{\rho} - \frac{\zeta'}{\zeta}(0) - \frac{1}{2}\log(1 - x^{-2})$$

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 Unfortuantely here there is a form of the Heisenberg uncertainty principle - to obtain finer detail of a function we need a wider range of information from the Fourier transform.

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Zero Density Estimates

Countin Zeros • More precisely to control this we need to truncate at height *T*.

$$\psi(x) = x - \sum_{\substack{\rho \\ |\gamma| \le T}} \frac{x^{\rho}}{\rho} + O\big(\log x + xT^{-1}\log^2(xT)\big).$$

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• If you study the proof of this in Chapter 12 you will see that this is about the best that one can hope to do in the current state of knowledge about how some things might be cancelling.

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- If you study the proof of this in Chapter 12 you will see that this is about the best that one can hope to do in the current state of knowledge about how some things might be cancelling.
- and so when $h \ll x$ we have

$$\psi(x+h) - \psi(x) = h - \sum_{\substack{\rho \\ |\gamma| \le T}} \frac{(x+h)^{\rho} - x^{\rho}}{\rho} + O\left(\log x + xT^{-1}\log^2(xT)\right)$$

• Clearly we will need to take T just a bit larger than $x(\log x)^2/h$. This emphasizes that the shorter h is longer T has to be.

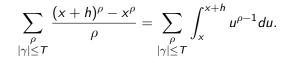
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• The crucial part is



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$$\sum_{\substack{\rho\\|\gamma|\leq T}}\frac{(x+h)^{\rho}-x^{\rho}}{\rho}=\sum_{\substack{\rho\\|\gamma|\leq T}}\int_{x}^{x+h}u^{\rho-1}du.$$

On RH the integral is ≪ hx^{-1/2}, so we certainly need to understand N(T), the number of zeros ρ with 0 < γ ≤ T. Note that by symmetry and the functional equation ρ is a zero of ζ iff any one of 1 − ρ, ρ and 1 − ρ is.

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- Thus on RH the total contribution from the zeros is

$$\ll h x^{-1/2} N(T) \ll h x^{-1/2} T \log T$$

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- Thus on RH the total contribution from the zeros is

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and we get

$$\psi(x+h) - \psi(x) \sim h$$

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for h a bit larger than $x^{1/2}(\log x)^3$.

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Countin Zeros • So, what if we do not assume RH?

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Counting Zeros • So, what if we do not assume RH? • In $\sum_{\substack{\rho \\ |\gamma| \le T}} \int_{x}^{x+h} u^{\rho-1} du$ we have $\int_{x}^{x+h} u^{\beta+i\gamma-1} du = x^{\beta+i\gamma-1} \int_{0}^{h} (1+v/x)^{\beta+i\gamma-1} dv.$

As v/x is small, the integrand does not vary much, and we don't know enough about the γ to get cancellation from $x^{i\gamma}$.

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 - As v/x is small, the integrand does not vary much, and we don't know enough about the γ to get cancellation from $x^{i\gamma}$.
- Thus in practice we have to deal with $\sum_{0 < \gamma \le T} hx^{\beta-1}$.
- We can write this as $hx^{-1} \sum_{0 < \gamma \le T} \left(x^{1/2} + \int_{1/2}^{\beta} x^u \log x du \right).$

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 - As v/x is small, the integrand does not vary much, and we don't know enough about the γ to get cancellation from $x^{i\gamma}$.
- Thus in practice we have to deal with $\sum_{0 < \gamma \leq T} h x^{\beta 1}$.
- We can write this as $hx^{-1} \sum_{0 < \gamma \le T} \left(x^{1/2} + \int_{1/2}^{\beta} x^u \log x du \right).$
- The $x^{1/2}$ we can treat as on RH. That leaves

$$hx^{-1} \sum_{0 < \gamma \le T} \int_{1/2}^{\beta} x^{u} \log x du = hx^{-1} \int_{1/2}^{1} N(u, T) x^{u} \log x du$$

where $N(u, T) = \operatorname{card}\{\rho : \beta \ge u, 0 \le \gamma \le T\}$.

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Zero Density Estimates

Countin Zeros A bound for N(u, T) = card{ρ : β ≥ u, 0 < γT} is known as a zero-density estimate.

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- A bound for N(u, T) = card{ρ : β ≥ u, 0 < γT} is known as a zero-density estimate.
- Well N(1, T) we know to be 0 and N(1/2, T) is about T log T, so the simple hypothesis is that

$$N(u,T) \ll T^{2(1-u)}(\log T)^B$$

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for some constant *B*. This is known as the *density hypothesis*.

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• If we just plug this in we run in to a problem. Actually let me assume only something a bit more general,

$$N(u,T) \ll T^{A(1-u)}(\log T)^B,$$

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since we know this for some choices of $A \ge 2$ and B.

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$$N(u,T) \ll T^{A(1-u)}(\log T)^B,$$

since we know this for some choices of $A \ge 2$ and B.

• We have reduced the sum over zeros to

$$hx^{-1} \int_{1/2}^{1} N(u, T) x^{u} \log x du$$
$$\ll hx^{-1} \int_{1/2}^{1} T^{A(1-u)} (\log T)^{B} x^{u} \log x du.$$

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•
$$\ll hx^{-1} \int_{1/2}^{1} T^{A(1-u)} (\log T)^B x^u \log x du$$
 and we suppose
 $T = x (\log x)^2 h^{-1}.$

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 and we suppose
 $T = x (\log x)^2 h^{-1}.$

Then this is

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$$\ll (x/h)^{A-1} (\log x)^{B'} \int_{1/2}^{1} (h^A x^{1-A})^u du$$

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and for h a bigger than $x^{1-1/A}$ the integral will be dominated by its value at 1.

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But this only gives

$$\ll h(\log x)^{B''}$$

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for some positive B''. Worse than trivial!

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If h is smaller than x^{1-1/A} the integral is dominated by its value at 1/2 but we only get the same sort of bound.

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and for h a bigger than $x^{1-1/A}$ the integral will be dominated by its value at 1.

• But this only gives

$$\ll h(\log x)^{B''}$$

for some positive B''. Worse than trivial!

- If *h* is smaller than $x^{1-1/A}$ the integral is dominated by its value at 1/2 but we only get the same sort of bound.
- Obviously we have to put in information about zeros near the 1-line.

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Motivation

Zero Density Estimates

Countin Zeros • If we insert the classical zero free region we obtain

$$hx^{-1} \int_{1/2}^{1-C/\log T} T^{A(1-u)} (\log T)^B x^u \log x du$$

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• which integrates to

$$(x/h)^{A-1}(h^A x^{1-A})^{1-C/\log T}(\log x)^{B''}$$

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- which is still not good enough since h^Ax^{1-A} and T are both some positive power of x.
- Until about 1938 the best zero-free region that was known was due to Littlewood, namely that all zeros of zeta satisfy

 $\beta \leq 1 - C \log \log(3 + |\gamma|) / \log(3 + |\gamma|).$

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• The extra loglog gives a bound

$$(x/h)^{A-1}(h^A x^{1-A})^{1-C(\log\log T)/\log T}(\log x)^{B''} = h(\log x)^{B''}(\log T)^{-\frac{C\log(h^A x^{1-A})}{\log T}}$$

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$$(x/h)^{A-1}(h^A x^{1-A})^{1-C(\log \log T)/\log T}(\log x)^{B''}$$
$$= h(\log x)^{B''}(\log T)^{-\frac{C\log(h^A x^{1-A})}{\log T}}$$

• log T should be of order log x. Since B" is rather large and C is rather small, in order to get a result

$$\frac{\log(h^A x^{1-A})}{\log T}$$

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will have to be largish. Substituting $h = x^{1-\delta}$ and T = x/h gives $(1 - \delta A)/\delta$ so δ has to be small.

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will have to be largish. Substituting $h = x^{1-\delta}$ and

- T = x/h gives $(1 \delta A)/\delta$ so δ has to be small.
- Thus Hoheisel needed to take $h = x^{\theta}$ with

$$\theta > 1 - \frac{1}{33000}.$$

[The other constraint $\theta > 1 - 1/A$ is far less demanding.]

• By working very hard Heilbronn reduced this to

$$\theta > 1 - \frac{1}{250}.$$

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$$eta \leq 1 - C \log^{\phi}(3 + |\gamma|).$$

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• If you plug this in you see that $(h^A x^{1-A})^{-C(\log T)^{-\phi}}$ will always kill any log power provided that $h > x^{\theta}$ where $\theta > 1 - 1/A$.

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- This leads to

$$\psi(x+h) - \psi(x) = h + O(h \exp(-c_2(\log x)^{\delta}))$$

for some positive constants c_2 and δ , provided that $h > x^{\theta}$ where $\theta > 1 - 1/A$.

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 To summarise. The fundamental ingredients are the explicit formula, an improved zero-free region and a zero density estimate of the form (where A ≥ 2)

$$N(u,T) \ll T^{A(1-u)} (\log T)^B$$

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 To summarise. The fundamental ingredients are the explicit formula, an improved zero-free region and a zero density estimate of the form (where A ≥ 2)

$$N(u,T) \ll T^{A(1-u)} (\log T)^B$$

• Note that we could replace this by $T \log T$ when $u \le 1 - 1/A$.

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Countin Zeros • Let me defer saying anything about N(u, T) and advert to some things connected with the explicit formula and the improved zero-free region.

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- Let me defer saying anything about N(u, T) and advert to some things connected with the explicit formula and the improved zero-free region.
- There is a general principle that the density of zeros of an entire function is related to the growth of the function.

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- Let me defer saying anything about N(u, T) and advert to some things connected with the explicit formula and the improved zero-free region.
- There is a general principle that the density of zeros of an entire function is related to the growth of the function.
- Thus the breakthrough in the proof of the explicit formula and the prime number theorem was Hadamard's beautiful theory of entire functions of bounded order.

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- In particular the ξ function is an entire function of order 1 and so

$$\zeta(s)(s-1)=e^{A+Bs}\prod_{
ho}(1-s/
ho)e^{s/
ho}$$

where here, and only here, the product is over all the zeros of $\boldsymbol{\zeta}.$

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- In particular the ξ function is an entire function of order 1 and so

$$\zeta(s)(s-1)=e^{\mathcal{A}+\mathcal{B}s}\prod_
ho(1-s/
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where here, and only here, the product is over all the zeros of $\boldsymbol{\zeta}.$

• There is a local form of this due to Landau which is commonly used and is recounted in Chapter 6 of MNT.

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Motivation

Zero Density Estimates

Counting Zeros

- Let me defer saying anything about N(u, T) and advert to some things connected with the explicit formula and the improved zero-free region.
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- There is a local form of this due to Landau which is commonly used and is recounted in Chapter 6 of MNT.
- It is also fundamental to the zero-free region.
- Thus it is important to have an estimate for the size of $\zeta(\sigma + it)$ when σ is close to 1 and t is large.

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Motivation

Zero Densit Estimates

Counting Zeros • To get the Chudakove style zero free region we need a bound of the kind

$$\zeta(\sigma+it) \ll \left(\log(2+|t|)\right)^{ heta}(2+|t|)^{(1-\sigma)^{1/ heta}}$$

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for some heta < 1, at least when $1 - \sigma \ll ig(\log(2+|t|)ig)^{- heta}$

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and by partial summation it suffices to treat sums

$$\sum_{n=N+1}^{N+M} n^{it} = \sum_{n=1}^{M} e^{it\log(N+n)} = N^{it} \sum_{n=1}^{M} \exp\left(it\log(1+n/N)\right)$$

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Motivation

Zero Densit Estimates

Countin Zeros • Now one can approximate log(1 + n/N) by a polynomial

$$\sum_{k=1}^{K} (-1)^{k-1} n^k N^{-k}.$$

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Motivation

Zero Density Estimates

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• Thus our sum becomes an exponential sum of the form

$$\sum_{n=1}^{M} e\left(\sum_{k=1}^{K} \alpha_k n^k\right).$$

Note that for reasons that only become apparent later we use $e(z) = e^{2\pi i z}$.

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• A standard way of dealing with such is to perturb it

$$\sum_{n=1}^{M} e\left(\sum_{k=1}^{K} \alpha_{k} n^{k}\right) = \sum_{n=r+1}^{M+r} e\left(\sum_{k=1}^{K} \alpha_{k} n^{k}\right) + O(r)$$
$$= \sum_{n=1}^{M} e\left(\sum_{k=1}^{K} \alpha_{k} (n+r)^{k}\right) + O(r)$$

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Zero Densit Estimates

Countin Zeros

• $\sum_{n=1}^{M} e\left(\sum_{k=1}^{K} \alpha_k (n+r)^k\right) + O(r).$

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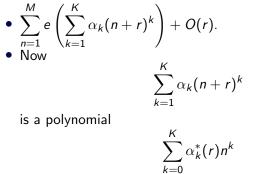
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Motivation

Zero Densit Estimates

Counting Zeros



in *n* and the coefficients α_k^* vary as *r* varies.

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Motivation

Zero Densit Estimates

Counting Zeros • $\sum_{\substack{n=1\\n=1}}^{M} e\left(\sum_{k=1}^{K} \alpha_k (n+r)^k\right) + O(r).$ • Now $\sum_{\substack{k=1\\n=1}}^{K} \alpha_k (n+r)^k$

is a polynomial

$$\sum_{k=0}^{K} \alpha_k^*(r) n^k$$

in *n* and the coefficients α_k^* vary as *r* varies.

• Then one can average over the r and look at

$$\frac{1}{R}\sum_{r=1}^{R}\left|\sum_{n=1}^{M}e\left(\sum_{k=1}^{K}\alpha_{k}^{*}(r)n^{k}\right)\right|^{2b}$$

• Now we can hope to relate this to

$$\int_{[0,1]^k} \left| \sum_{n=1}^M e\left(\sum_{k=1}^K x_k n^k \right) \right|^{2b} dx_1 \dots dx_K.$$

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Zero Densit Estimates

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• $\int_{[0,1]^k} \left| \sum_{n=1}^M e\left(\sum_{k=1}^K x_k n^k \right) \right|^{2b} dx_1 \dots dx_K.$

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Motivation

Zero Densit Estimates

Counting Zeros

$$\int_{[0,1]^k} \left| \sum_{n=1}^M e\left(\sum_{k=1}^K x_k n^k \right) \right|^{2b} dx_1 \dots dx_K.$$

• By the orthogonality relationship

$$\int_0^1 e(xh)dx = \begin{cases} 1 & h = 0, \\ 0 & h \in \mathbb{Z} \setminus \{0\}, \end{cases}$$

this is the number of solutions of

$$m_1 + \dots + m_b = n_1 + \dots + n_b$$

$$\vdots$$

$$m_1^k + \dots + m_b^k = n_1^k + \dots + n_b^k$$

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 $m_1^K + \cdots + m_b^K = n_1^K + \cdots + n_b^K.$

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Motivation

Zero Densit<u>y</u> Estimates

Counting Zeros

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$$m_1^K + \cdots + m_b^K = n_1^K + \cdots + n_b^K.$$

 Any non-trivial estimate for this is known as the "Vinogradov Mean Value Theorem" and this is the core part of Chapter 24.

> Robert C. Vaughan

Motivation

Zero Density Estimates

Countin Zeros • Now let me return bounds for N(u, T).

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Motivation

Zero Density Estimates

Countin Zeros

- Now let me return bounds for N(u, T).
- Ingham in 1937 had obtained the bound

$$N(u,T) \ll T^{rac{3(1-u)}{2-u}}(\log T)^5$$

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improving on Carlson, Landau, Titchmarsh and Hoheisel.

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• Since for $\frac{1}{2} \le u < 1$ we have

$$\frac{3}{2-u} \le 3$$

it follows that when
$$x = h^{\theta}$$
, $\theta > \frac{2}{3}$,

$$\psi(x+h) - \psi(x) \sim h$$

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• Montgomery in 1969 obtained

$$N(u,T) \ll T^{\frac{2(1-u)}{u}} (\log T)^{14}$$

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which improves on lngham when $u > \frac{4}{5}$.

• Moreover they collectively give

$$N(u, T) \ll T^{\frac{8(1-u)}{5}} (\log T)^{14}$$

and so we get the desired estimate when $\theta > \frac{5}{8}$.

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Motivation

Zero Density Estimates

Counting Zeros • In 1972 Huxley refined Montgomery's method to give

$$N(u, T) \ll T^{\frac{(5u-3)(1-u)}{u^2+u-1}} (\log T)^9$$

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and this is better than Ingham when u > 3/4• In particular for all $u \in [1/2, 0)$

$$N(u,T) \ll T^{\frac{12(1-u)}{5}} (\log T)^9$$

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$$N(u,T) \ll T^{rac{3(1-u)}{3u-1}} (\log T)^C$$

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• Important Remark. These exponents all give an $A < \frac{12}{5}$ when u is not close to $\frac{3}{4}$. So you only have to improve the zero density estimates in the neighbourhood of $u = \frac{3}{4}$. This is what Guth and Maynard do. Thus to get the final result you also need to know the earlier work!

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Motivation

Zero Densit Estimates

Counting Zeros • So how does one prove a zero-density estimate?

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Motivation

Zero Density Estimates

Counting Zeros

- So how does one prove a zero-density estimate?
- Let me outline Montgomery's proof of Ingham's bound. The details are in Chapter 28.

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- An idea which has numerous important applications in analytic number theory is the observation that

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 when $1 < n \leq K$.

• Then for $\sigma > 1$

$$1+\sum_{n>K}a(n)n^{-s}=M(s)\zeta(s)$$

where
$$M(s) = \sum_{k=1}^{K} \mu(k) k^{-s}$$
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 In that half-plane M(s)ζ(s) is close to 1 if K is quite large and we might hope that persists when 1/2 < Re s = σ < 1. Indeed this can be shown to be true on RH.

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- In that half-plane M(s)ζ(s) is close to 1 if K is quite large and we might hope that persists when 1/2 < Re s = σ < 1. Indeed this can be shown to be true on RH.
- On the other hand $M(s)\zeta(s)$ will be 0 when $\zeta(s) = 0$.

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Motivation

Zero Densit Estimates

Counting Zeros • One way this can be realised is through the transformation $\sum_{n=1}^{\infty} a(n)n^{-s}e^{-n/Y}$ $= \frac{1}{2\pi i} \int_{2-i\infty}^{2+i\infty} M(s+w)\zeta(s+w)Y^w\Gamma(w)dw.$

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- The integrand has singularities at 0, and w = 1 s.
- In view of the bounds given by Corollaries 10.5 and 10.10, Lemma 10.15 and (C.19) of Theorem C.1 we are able to move the path of integration to the line Re $w = \frac{1}{2} - \sigma$ and pick up the residues at 0 and w = 1 - s.

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- Thus

$$\sum_{n=1}^{\infty} a(n)n^{-s}e^{-n/Y} =$$

$$M(1)Y^{1-s}\Gamma(1-s) + M(s)\zeta(s) +$$

$$\int_{-\infty}^{\infty} M(\frac{1}{2} + it + iv)\zeta(\frac{1}{2} + it + iv)Y^{\frac{1}{2} - \sigma + iv}\Gamma(\frac{1}{2} - \sigma + iv)\frac{dv}{2\pi}.$$

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Motivation

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$$\sum_{n=1}^{\infty} a(n)n^{-s}e^{-n/Y} = M(1)Y^{1-s}\Gamma(1-s) + M(s)\zeta(s) + \int_{-\infty}^{\infty} M(\frac{1}{2} + it + iv)\zeta(\frac{1}{2} + it + iv)Y^{\frac{1}{2} - \sigma + iv}\Gamma(\frac{1}{2} - \sigma + iv)\frac{dv}{2\pi}.$$

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• If
$$\zeta(\rho) = 0$$
 we have

$$e^{-\frac{1}{Y}} + \sum_{n=K+1}^{\infty} a(n)n^{-\rho}e^{-\frac{n}{Y}} = M(1)Y^{1-\rho}\Gamma(1-\rho) + \int_{-\infty}^{\infty} M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)Y^{\frac{1}{2}-\beta+i\nu}\Gamma(\frac{1}{2}-\beta+i\nu)\frac{d\nu}{2\pi}.$$

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Motivation

Zero Densit Estimates

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•
$$\sum_{n=1}^{\infty} a(n)n^{-s}e^{-n/Y} = M(1)Y^{1-s}\Gamma(1-s) + M(s)\zeta(s) + \int_{-\infty}^{\infty} M(\frac{1}{2} + it + iv)\zeta(\frac{1}{2} + it + iv)Y^{\frac{1}{2} - \sigma + iv}\Gamma(\frac{1}{2} - \sigma + iv)\frac{dv}{2\pi}.$$

• If $\zeta(\rho) = 0$ we have

$$e^{-\frac{1}{Y}} + \sum_{n=K+1}^{\infty} a(n)n^{-\rho}e^{-\frac{n}{Y}} = M(1)Y^{1-\rho}\Gamma(1-\rho) + \int_{-\infty}^{\infty} M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)Y^{\frac{1}{2}-\beta+i\nu}\Gamma(\frac{1}{2}-\beta+i\nu)\frac{d\nu}{2\pi}.$$

• We have $\Gamma(1-\rho) \ll e^{-|\gamma|}$, so for $K = T \leq Y \ll T^2$, $Z = Y(\log T)^2$ and $(\log T)^2 \leq \gamma \leq T$ have

$$e^{-1/Y} + \sum_{K < n \leq Z} a(n) n^{-
ho} e^{-n/Y} = O(T^{-1}) +$$

 $\int_{-\infty}^{\infty} M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)Y^{\frac{1}{2}-\beta+i\nu}\Gamma(\frac{1}{2}-\beta+i\nu)\frac{d\nu}{2\pi}.$

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Motivation

Zero Densit Estimates

Counting Zeros

•
$$e^{-1/Y} + \sum_{K < n \le Z} a(n) n^{-\rho} e^{-n/Y} = O(T^{-1}) +$$

$$\int_{-\infty}^{\infty} M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)Y^{\frac{1}{2}-\beta+i\nu}\Gamma(\frac{1}{2}-\beta+i\nu)\frac{d\nu}{2\pi}.$$

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Counting Zeros

$$e^{-1/Y} + \sum_{K < n \le Z} a(n) n^{-\rho} e^{-n/Y} = O(T^{-1}) + O(T^{-1})$$

$$\int_{-\infty}^{\infty} M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)Y^{\frac{1}{2}-\beta+i\nu}\Gamma(\frac{1}{2}-\beta+i\nu)\frac{d\nu}{2\pi}.$$

• Thus

T

$$1 \ll \left| \sum_{K < n \le Z} a(n) n^{-\rho} e^{-n/Y} \right| + Y^{\frac{1}{2} - \beta} \int_{-\infty}^{\infty} |M(\frac{1}{2} + i\gamma + i\nu)\zeta(\frac{1}{2} + i\gamma + i\nu)|e^{-|\nu|}d\nu.$$

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Counting Zeros

$$e^{-1/Y} + \sum_{K < n \le Z} a(n) n^{-\rho} e^{-n/Y} = O(T^{-1}) + O(T^{-1})$$

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$$1 \ll \left| \sum_{K < n \leq Z} a(n) n^{-\rho} e^{-n/Y} \right| + Y^{\frac{1}{2} - \beta} \int_{-\infty}^{\infty} |M(\frac{1}{2} + i\gamma + i\nu)\zeta(\frac{1}{2} + i\gamma + i\nu)|e^{-|\nu|}d\nu.$$

• Hence

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Motivation

Zero Densit Estimates

Counting Zeros

$$1 \ll \Big| \sum_{K < n \leq Z} a(n) n^{-\rho} e^{-n/Y} \Big| +$$

$$Y^{\frac{1}{2}-\beta}\int_{-\infty}^{\infty}|M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)|e^{-|\nu|}d\nu.$$

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Motivation

Zero Densit Estimates

Counting Zeros

$$\ll \left| \sum_{K < n \leq Z} a(n) n^{-\rho} e^{-n/Y} \right| +$$
$$Y^{\frac{1}{2} - \beta} \int_{-\infty}^{\infty} |M(\frac{1}{2} + i\gamma + iv)\zeta(\frac{1}{2} + i\gamma + iv)| e^{-|v|} dv.$$

• We can partition the set \mathcal{R} of zeros counted by N(u, T)into two classes \mathcal{R}_1 and \mathcal{R}_2 , in the first of which the first sum is larger and the second contains the remainder, and let their counts be $N_1(u, T)$ and $N_2(u, T)$. This gives tremendous flexibility.

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Motivation

Zero Density Estimates

Counting Zeros

$$\ll \Big| \sum_{K < n \leq Z} a(n) n^{-\rho} e^{-n/Y} \Big| +$$
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- We can partition the set \mathcal{R} of zeros counted by N(u, T)into two classes \mathcal{R}_1 and \mathcal{R}_2 , in the first of which the first sum is larger and the second contains the remainder, and let their counts be $N_1(u, T)$ and $N_2(u, T)$. This gives tremendous flexibility.
- In the proof of Ingham's bound we use

$$N_1(u, T) \ll \sum_{\rho \in \mathcal{R}} \Big| \sum_{K < n \leq Z} a(n) n^{-\rho} e^{-n/Y} \Big|^2$$

and $N_2(u, T) \ll$ $Y^{\frac{2}{3}-\frac{4u}{3}} \int_{-\infty}^{\infty} \sum_{\rho \in \mathcal{R}} \left| M(\frac{1}{2}+i\gamma+iv) \zeta(\frac{1}{2}+i\gamma+iv) \right|^{\frac{4}{3}} e^{-|v|} dv.$

This makes best use of the various mean value theorems we have for Dirichlet polynomials.

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Motivation

Zero Densit<u>:</u> Estimates

Counting Zeros • We already saw in the outline of the Chudakov-Korobov-Vinogradov zero free region that one needs to relate discrete mean values to continuous mean values and we will see more of this here.

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Motivation

Zero Density Estimates

Counting Zeros

- We already saw in the outline of the Chudakov-Korobov-Vinogradov zero free region that one needs to relate discrete mean values to continuous mean values and we will see more of this here.
- Those who have seen the large sieve, say in Math 571 or Math 572 will be aware of the inequality

$$\sum_{r=1}^{R} |S(x_r)|^2 \ll (N+\delta^{-1}) \int_0^1 |S(x)|^2 dx$$

where

$$S(x) = \sum_{m=1M+1}^{M+N} c_n e(nx)$$

and

$$||x_q - x_r|| := \min_{n \in \mathbb{Z}} |x_q - x_r - n| \ge \delta$$

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and

$$\|x_q - x_r\| := \min_{n \in \mathbb{Z}} |x_q - x_r - n| \ge \delta.$$

 A method of Gallagher which will give this can be adapted to many situations, such as sums over zeros ρ.

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Motivation

Zero Densit Estimates

Counting Zeros

• $f(x) - f(y) = \int_{y}^{x} f'(v) dv$

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Motivation

Zero Densit Estimates

Counting Zeros

•
$$f(x) - f(y) = \int_{y}^{x} f'(v) dv$$

• $\int_{x-\delta}^{x+\delta} (f(x) - f(y)) dy = \int_{x-\delta}^{x+\delta} \int Y^{x} f'(v) dv dy$

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Motivation

Zero Densit Estimates

Counting Zeros

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• $\int_{x-\delta}^{x+\delta} (f(x) - f(y)) dy = \int_{x-\delta}^{x+\delta} \int Y^{x} f'(v) dv dy$

• This can be rearranged to give

$$2\delta f(x) = \int_{x-\delta}^{x+\delta} f(y) dy + \int_{x-\delta}^{x+\delta} f'(v) sgn(x-v)(\delta - |x-v|) dv$$

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Motivation

Zero Densit Estimates

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•
$$|f(x)| \leq rac{1}{2\delta} \int_{x-\delta}^{x+\delta} |f(y)| dy + rac{1}{2} \int_{x-\delta}^{x+\delta} |f'(v)| dv$$

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Motivation

Zero Densit Estimates

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$$f(x) - f(y) = \int_{y}^{x} f'(v) dv$$

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•
$$|f(x)| \leq rac{1}{2\delta}\int_{x-\delta}^{x+\delta}|f(y)|dy+rac{1}{2}\int_{x-\delta}^{x+\delta}|f'(v)|dv$$

• Say something about spacing and two dimensions.

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Motivation

Zero Densit Estimates

Counting Zeros • To return to Ingham.

$$N_1(u, T) \ll \sum_{
ho \in \mathcal{R}} \Big| \sum_{K < n \leq Z} a(n) n^{-
ho} e^{-n/Y} \Big|^2$$

and $N_2(u, T) \ll$

$$Y^{\frac{2}{3}-\frac{4u}{3}}\int_{-\infty}^{\infty}\sum_{\rho\in\mathcal{R}}\left|M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)\right|^{\frac{4}{3}}e^{-|\nu|}d\nu.$$

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The first inequality will give N₁(u, T) ≪ Y^{2(1-u)}(log T)^C, which looks promising, especially if we could take Y ≈ T.

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Motivation

Zero Densit Estimates

Counting Zeros • To return to Ingham.

$$N_1(u, T) \ll \sum_{\rho \in \mathcal{R}} \Big| \sum_{K < n \leq Z} a(n) n^{-\rho} e^{-n/Y} \Big|^2$$

and $N_2(u, T) \ll$

$$Y^{\frac{2}{3}-\frac{4u}{3}}\int_{-\infty}^{\infty}\sum_{\rho\in\mathcal{R}}\left|M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)\right|^{\frac{4}{3}}e^{-|\nu|}d\nu.$$

- The first inequality will give N₁(u, T) ≪ Y^{2(1-u)}(log T)^C, which looks promising, especially if we could take Y ≈ T.
- The second inequality is trickier. If we knew the mean-value version of the Lindelöf hypothesis ζ(1/2 + it) ≪_ε (2 + |t|)^ε, namely

$$\int_0^T |\zeta(1/2+it)|^k dt \ll_{\varepsilon,k} T^{1+\varepsilon},$$

then we could also deduce $N_2(u, T) \ll T^{2(1-u)+\varepsilon}$. Thus Lindelöf \Rightarrow MVLindelöf \Rightarrow Density Hypothesis $\Rightarrow \theta > \frac{1}{2}/2$.

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Motivation

Zero Density Estimates

Counting Zeros • Unfortunately the best mean value theorem we have for ζ on the $\frac{1}{2}\text{-line}$ is

$$\int_0^T |\zeta(1/2+it)|^4 dt \ll T(\log T)^4.$$

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Motivation

Zero Density Estimates

Counting Zeros • Unfortunately the best mean value theorem we have for ζ on the $\frac{1}{2}\text{-line}$ is

$$\int_0^T |\zeta(1/2+it)|^4 dt \ll T(\log T)^4.$$

• This motivates the $rac{4}{3}$ in $N_2(u,T)\ll$

$$Y^{\frac{2}{3}-\frac{4u}{3}}\int_{-\infty}^{\infty}\sum_{\rho\in\mathcal{R}}\left|M(\frac{1}{2}+i\gamma+i\nu)\zeta(\frac{1}{2}+i\gamma+i\nu)\right|^{\frac{4}{3}}e^{-|\nu|}d\nu.$$

since we can use Hölder's inequality to give

$$\int_{0}^{T} |M(\frac{1}{2} + it)\zeta(\frac{1}{2} + it)|^{\frac{4}{3}} dt \leq \\ \left(\int_{0}^{T} |M(\frac{1}{2} + it)|^{2} dt\right)^{\frac{2}{3}} \left(\int_{0}^{T} |\zeta(\frac{1}{2} + it)|^{4} dt\right)^{\frac{1}{3}}.$$

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Motivation

Zero Densit<u>y</u> Estimates

Counting Zeros • The Huxley proof is somewhat more sophisticated and involves in addition a version of mean value theorems called a large values theorem. This is where the breakthrough occurs in the Guth-Maynard paper.

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